





Causal inference and causal discovery with perspectives in Earth sciences

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Lecture developed together with Andreas Gerhardus

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DLR Institute of Data Science

Knowledge for Tomorrow



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Motivation







Consider the following setup:

- You have a lawn with a sprinkler.
- If the sprinkler is on, the lawn will become wet.
- Each day you toss a coin and turn the sprinkler on if it shows heads.



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Two easy questions:

- Does turning on the sprinkler cause the lawn to be wet?
- Does making the lawn wet turn the sprinkler on?



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Two easy questions:

Does turning on the sprinkler cause the lawn to be wet?

Yes!

Does making the lawn wet turn the sprinkler on?

No!





How does this relate to statistical dependence?

We can model the setup with two random variables:

X: X = 1 if the sprinkler is on, else X = 0

Y: Y = 1 if the lawn is wet, else Y = 0



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We can model the setup with two random variables:

$$X: X = 1$$
 if the sprinkler is on, else $X = 0$

$$Y: Y = 1$$
 if the lawn is wet, else $Y = 0$

Since turning the sprinkler on causes the lawn to be wet, we get

$$\mathbb{P}(\text{lawn wet}|\text{sprinkler on}) = \mathbb{P}(Y = 1|X = 1)$$

$$> \mathbb{P}(Y=1) = \mathbb{P}(\mathsf{lawn} \ \mathsf{wet})$$
 .

This shows that X and Y are dependent random variables.



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This shows that *X* and *Y* are dependent random variables.

However, the dependence also goes in the other direction. Observing the lawn to be wet increases the probability that the sprinkler is on:

$$\mathbb{P}(\mathsf{sprinkler}\;\mathsf{on}|\mathsf{lawn}\;\mathsf{wet}) = \mathbb{P}(X=1|Y=1)$$

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The probabilities alone do not allow us to decide wheter X causes







Conceptual Findings

- Statistical dependence is symmetric between cause and effect, i.e., it does not distinguish between cause and effect
- Causation is asymmetric between cause and effect, i.e., it does distinguish between cause and effect
- Statistical dependence is not causation.



Observational Causal Inference







Causal Inference

What causal inference is about:

- Provide mathematical language for causal notions such as cause and effect
- Provide methods and assumptions for answering causal questions from observational data
- Provide methods and assumptions for learning parts of the causal relationships from observational data





Question:

How can one, in general, determine whether a certain drug is helpful in reducing the risk of developing a certain desease?



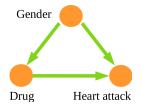


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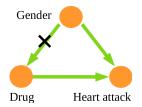


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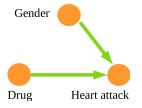


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The gender of a subject does not longer influence whether a subject belongs to the treatment or control group.



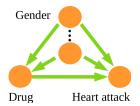


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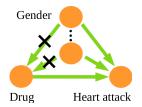


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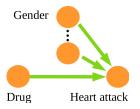


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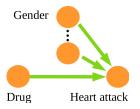


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In fact, all confounders are eliminated! Therefore, in RCT data causal effects can be assessed by statistical proportions.





Experimentation

Abstraction of the question:

Consider two events or variables X and Y. How can we determine whether X causes Y?

Experimentation:

Experimentally manipulate X while keeping all other conditions exactly the same. If this results in a change of Y, X causes Y.

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- Turning the sprinkler on makes the lawn wet.
- Making the lawn wet does not turn on the sprinkler.

We need to imagine the lawn being made wet without changing something else, in particular without turning the sprinkler on









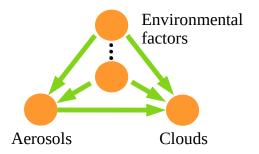


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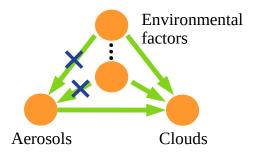






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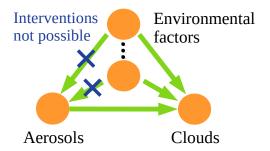






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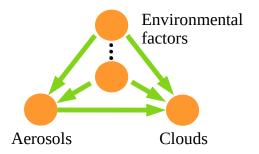






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For most parts it is impossible, and else arguably ethically questionable, to deliberately intervene on climate variables.

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Causal relationships are mostly known among microscopic physical variables, but usually less so among macroscopic variables.





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For example:

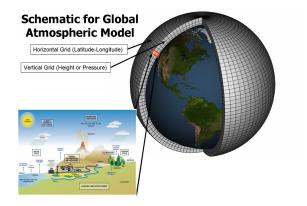
- Is there a causal relationship between arctic sea ice extent in winter and mid latitude weather in summer?
- Is there a causal relationship between ENSO and extreme weather events somewhere on Earth?





Simulation:

Experimentation in climate models: Fix (intervene on) certain parameters and analyze the resulting effect on other parameters.





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Seminal work by Judea Pearl, Peter Spirtes, Clark Glymour, Richard Scheines.















Simulation:

Experimentation in climate models: Fix (intervene on) certain parameters and analyze the resulting effect on other parameters.

Observational causal inference:

Answer causal questions, for example questions about cause and effect, from observation data.

To mathematically treat causal inference, we need to formalize the idea of an underlying structural causal model (SCM), define hypothetical interventions, and establish identifiability criteria to decide if and how causal effects can be estimated from data alone (without interventions).





Intuition:

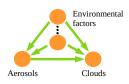
A structural causal model (SCM) specifies the functional causal relationships between a set of random variables.

Example:

$$X_{\text{aerosols}} := f_{\text{aerosols}}(X_{\text{env. facs.}}, \eta_{\text{aerosols}})$$

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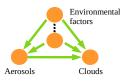
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Claims of an SCM:

- Specifies the *direct causes*, also called *parents*, of each variable
- The functions f_i are independent mechanisms by which nature determines the values of each variable based on the values of its direct causes plus random fluctuations





Definition:

A structural causal model over the set of random variables

 $\mathbf{X} = \{X_1, \dots, X_n\}$ consists auf n structural assignments

$$X_1 := f_1(PA_1, \eta_1)$$

$$\vdots$$

$$X_n := f_n(PA_n, \eta_n)$$

together with a specification of the probability distributions of the jointly independent 'noise' variables η_i . Here $PA_i \subseteq \mathbf{X} \setminus \{X_i\}$ are the direct causes of X_i .

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Interpretation of noise variables:

- Summarize all background factors outside causal model, i.e., all factors apart from X_1, \ldots, X_n
- Their joint independence means the model is sufficient to describe

 the causal relationship among the variables X_1, \dots, X_n



Causal Graphs

Definition:

Consider an SCM over the variables $\mathbf{X} = \{X_1, \dots, X_n\}$. Its causal graph \mathcal{G} is the directed graph with

- 1. a vertex (node) for each variable X_i
- 2. for all i a directed edge from each variable $X_i \in PA_i$ to X_i





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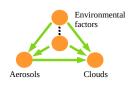
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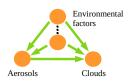
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Interpretation:

- Nodes represent variables
- Edges represent causal influences







Definition:

Consider SCM over the variables $\mathbf{X} = \{X_1, \dots, X_n\}$.

An intervention on X_i , denoted as $do(X_i := x_{i,0})$, defines a modified SCM obtained by replacing the old assignment of X_i with $X_i := x_{i,0}$ where $x_{i,0}$ is some value in the range of X_i :

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Interpretation:

- An intervention $do(X_i := x_{i,0})$ inactivates the natural mechanism $X_i := f_i(PA_i, \eta_i)$ by forcing X_i to take the value $x_{i,0}$
- All other mechanisms remain unmodified
- Interventions can also be more complex functions than just $x_{i,0}$



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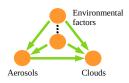
Example:

Original SCM prior to intervention:

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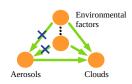
Example:

After intervention $do(X_{aerosols}) := a$ certain aerosol concentration:

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General formalism:

- Formally define the notion of a causal effect where generally $\mathbb{P}(Y|do(X:=x)) \neq \mathbb{P}(Y|X=x).$
- Based on the causal graph, decide whether the causal effect of X on Y can be identified from the observational data.
- If the effect can be identified, express causal effect in terms of observational data





General formalism:

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Here: Convey idea through examples.





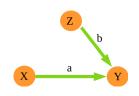
Example 1: No confounding

Consider the linear SCM:

$$Z := \eta_Z$$
, $\eta_Z \sim \mathcal{N}(0, \sigma_Z^2)$

$$X := \eta_X , \qquad \qquad \eta_X \sim \mathcal{N}(0, \sigma_X^2)$$

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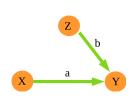
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Goal: Determine a, the causal effect of X on Y







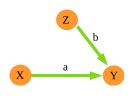
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Goal: Determine a, the causal effect of X on Y

Idea: Use covariance of Y on X

$$Cov(Y,X) = Cov(a \cdot \eta_X + b \cdot \eta_Z + \eta_Y, \eta_X,)$$

$$= a \cdot Cov(\eta_X, \eta_X) + b \cdot Cov(\eta_Z, \eta_X) + Cov(\eta_Y, \eta_Z)$$

$$= a \cdot Var(X) + b \cdot 0 + 0$$

$$\Rightarrow a = \frac{Cov(Y,X)}{Var(X)} = \rho(Y,X) \cdot \sqrt{\frac{Var(Y)}{Var(X)}}$$





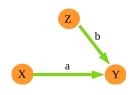
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$$\begin{aligned} \textit{Cov}\big(Y,X\big) &= \textit{Cov}\big(a \cdot \eta_X + b \cdot \eta_Z + \eta_Y, \eta_X,\big) \\ &= a \cdot \textit{Cov}\big(\eta_X, \eta_X\big) + b \cdot \textit{Cov}\big(\eta_Z, \eta_X\big) + \textit{Cov}\big(\eta_Y, \eta_Z\big) \\ &= a \cdot \textit{Var}(X) + b \cdot 0 + 0 \\ &\Rightarrow a = \frac{\textit{Cov}(Y,X)}{\textit{Var}(X)} = \rho(Y,X) \cdot \sqrt{\frac{\textit{Var}(Y)}{\textit{Var}(X)}} \end{aligned}$$



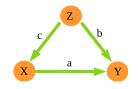
Example 2: Observed confounding

Consider the linear SCM:

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$$X := \mathbf{c} \cdot \mathbf{Z} + \eta_X , \qquad \qquad \eta_X \sim \mathcal{N}(0, \sigma_X^2)$$

$$Y := a \cdot X + b \cdot Z + \eta_Y \ , \ \ \ \eta_Y \sim \mathcal{N}(0, \sigma_Y^2)$$





Example 2: Observed confounding

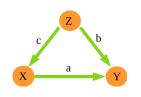
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 $\underline{\text{Goal:}}$ Determine a, the causal effect of X on Y







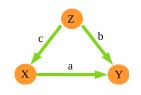
Example 2: Observed confounding

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Goal: Determine a, the causal effect of X on Y

First attempt: Linearly regress of Y on X

$$Cov(Y,X) = Cov((a \cdot c + b) \cdot \eta_Z + a \cdot \eta_X + \eta_Y, c \cdot \eta_Z + \eta_X)$$
$$= c \cdot (a \cdot c + b) \cdot Cov(\eta_Z, \eta_Z) + a \cdot Cov(\eta_X, \eta_X)$$

$$Var(X) = Cov(c \cdot \eta_Z + \eta_X, c \cdot \eta_Z + \eta_X)$$

= $c^2 \cdot Cov(\eta_Z, \eta_Z) + Cov(\eta_X, \eta_X)$





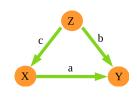




Example 2: Observed confounding

Second attempt:

- Control for Z: Regress out influence of Z on X and Y to obtain residuals ΔX and ΔY
- Linearly regress ΔY on ΔX





Example 2: Observed confounding

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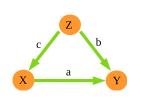
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- Linearly regress ΔY on ΔX

$$\Delta X = X - \frac{Cov(X, Z)}{Var(Z)} \cdot Z$$

$$= X - \frac{c \cdot Cov(\eta_Z, \eta_Z)}{Cov(\eta_Z, \eta_Z)} \cdot Z$$

$$= X - c \cdot Z$$

$$= \eta_X$$



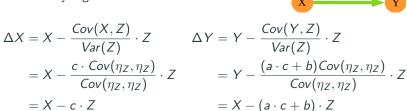




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 $= a \cdot \eta_X + \eta_Y$



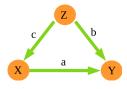
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$$\Delta X = X - \frac{Cov(X, Z)}{Var(Z)} \cdot Z \qquad \Delta Y = Y - \frac{Cov(Y, Z)}{Var(Z)} \cdot Z$$

$$= X - \frac{c \cdot Cov(\eta_Z, \eta_Z)}{Cov(\eta_Z, \eta_Z)} \cdot Z \qquad = Y - \frac{(a \cdot c + b)Cov(\eta_Z, \eta_Z)}{Cov(\eta_Z, \eta_Z)} \cdot Z$$

$$= X - c \cdot Z \qquad = X - (a \cdot c + b) \cdot Z$$

$$= \eta_X \qquad = a \cdot \eta_X + \eta_Y$$

$$\Rightarrow a = \frac{Cov(\Delta Y, \Delta X)}{Var(\Delta X, \Delta X)} = \frac{a \cdot Cov(\eta_X, \eta_X)}{Cov(\eta_X, \eta_X)} = a$$







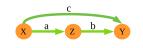
Example 3: Partial mediation

Consider the linear SCM:

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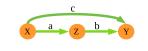
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Total causal effect of X on Y: $(a \cdot b + c)$

• Linearly regress Y on X

Direct causal effect of X on Y: c

• Multivariate regression of Y on its parents (X, Z), then take coefficient belonging to X





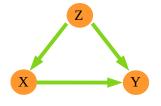
Conceptual Findings

- Interventions model the notion of 'experimentally manipulate X_i while keeping all other conditions exactly the same'.
- The causal graph allows us to decide whether a certain causal effect is identifiable.
- If the causal effect is identifiable, the causal graph further allows us to estimate it from observational data.



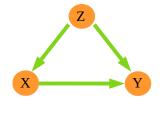


Example: Observed confounder





Example: Observed confounder



Control for confounder Z



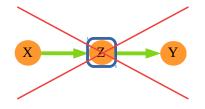
Example: Mediation







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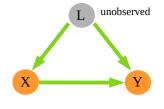


Do not control for mediator Z





Example: Unobserved confounder

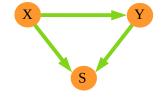


Causal effect cannot be identified



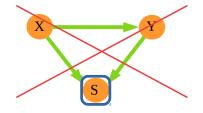


Example: Common effect



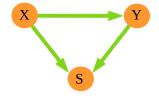


Example: Common effect





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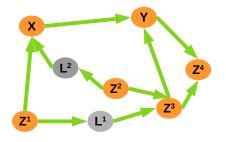
Do not control for common effects of X and Y, this would introduce bias.

However, sometimes the available dataset is already biased, e.g., if samples are missing (selection bias).





Example: Complicated case



Back-door criterion:

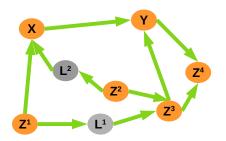
Block all non-causal dependence by controlling for an appropriate set of variables



Back-door criterion

A set $\mathbf{Z} \subset \mathbf{X} \setminus \{X, Y\}$ satisfies the back-door criterion for identifying the causal effect of X on Y if

1. **Z** blocks all paths between X and Y with arrow into X



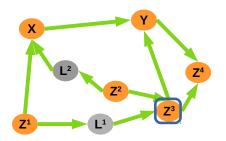




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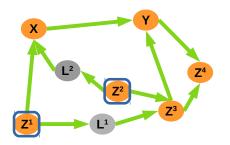




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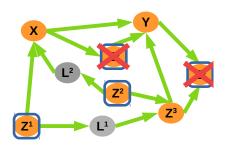




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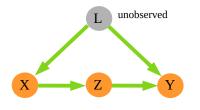
Theorem: Back-door adjustment as a sufficient criterion If Z satisfy the back-door criterion for (X, Y), then

$$p(y|\frac{do(x)}{}) = \sum_{\mathbf{z}} p(y|x,\mathbf{z}) \cdot p(\mathbf{z})$$





Example: Unobserved confounder plus mediation



Front-door criterion:

Allows to identify causal effect, but not as simple as controlling for a certain set of variables





Observational Causal Discovery







How to Get Causal Graphs

Relevance: Result of previous section

Knowing the causal graph of the data generating process, we can determine whether and how a causal effect can be identified from observational data



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Knowing the causal graph of the data generating process, we can determine whether and how a causal effect can be identified from observational data

How to get the causal graph?

Option 1: Scientific knowledge

Talk to domain experts, general reasoning

Option 2: Observational causal discovery

Learn from observational data





Fact:

The structure of the causal graph often has observable implications in terms of (conditional) independencies in the observed data.

Intuition:

- Statistical dependencies derive from causal relationships
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- X_1 influences X_2 :
- X₂ influences X₃:
- • X_1 influences X_3 through X_2 :
- Knowing X_2 , X_1 does not say more about X_3 : $X_1 \perp \!\!\! \perp X_3 \mid X_2$



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Example 2: Fork



- $X_1 \cancel{\perp} X_2$ $X_2 \cancel{\perp} X_3$ • X_2 influences X_1 :
- X_2 influences X_3 :
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Example 2: Fork



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Example 3: Collider



- $X_1 \cancel{\perp} X_2$ $X_2 \cancel{\perp} X_3$ • X_1 influences X_2 :
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- $X_1 \perp \!\!\! \perp X_3$ • No influence between X_1 and X_3 :
- Observing X_2 introduces selection bias between $X_1 \perp \!\!\! \perp X_3 \mid X_2$ X_1 and X_3 :





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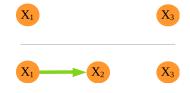
 $X_1 \perp \!\!\! \perp X_3$

 $X_1 \perp \!\!\! \perp X_3 \mid X_2$





Example 4: Disconnected variables



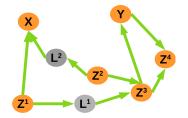
• No influence between X_1 and X_3 : $X_1 \perp \!\!\! \perp X_3$



General rule: d-separation

The graphical criterion of *d-separation* allows to read off all (conditional) independencies implied by the structure of a particular causal graph.

Example 5: Complicated case



Some independencies implied by d-separation

- $X \perp \!\!\!\perp Y|Z^1, Z^2$
- *X* ⊥⊥ *Y*|*Z*³







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Possible causal graphs:











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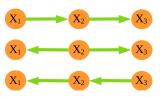
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observationally equivalent graphs







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X₁ ± X₂
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The data generating process can be described as an SCM. This implies the so called *causal Markov condition*:

d-separation in causal graph \Rightarrow (conditional) independence

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Assumption 3: Causal sufficiency

There are no unobserved confounders and there are no selection variables.

Assumption 3 can be dropped (but much more complicated)





Some Famous Algorithms

SGS-Algorithm:

Tests all possible (conditional) independence statements, use results to constrain the causal graph as discussed in the examples.

SGS stands for Spirtes, Glymour, and Scheines.

PC-Algorithm:

Based on SGS-Algorithm, but using much fewer conditional independence tests.

PC stands for Peter Spirtes and Clark Glymour.

FCI-Algorithm:

Similar to the PC-Algorithm, but without requiring the assumption of causal sufficiency, i.e., allowing for hidden confounders and selection variabes.

FCI stands for fast causal inference.





Two Main Approaches of Causal Discovery

Constraint-based causal discovery:

Constrain causal graph by using results of conditional independence tests in observational data.

⇒ discussed today

SCM-based causal discovery:

Make assumption on functional causal relationships (e.g., linear or non-linear) and noise distributions (e.g., Gaussian or non-Gaussian) of data generating SCM.

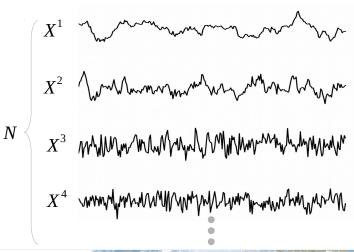
Generically, model can fit in one direction only. This allows to identify direction of causal influence.

 \Rightarrow not discussed today





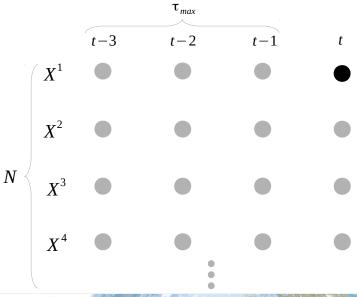
Causal Discovery in Time Series







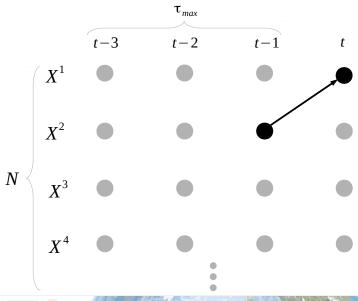
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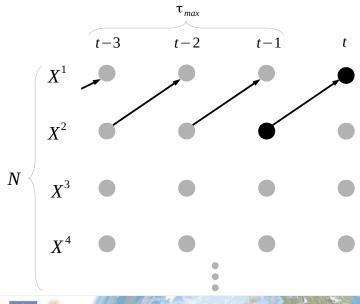


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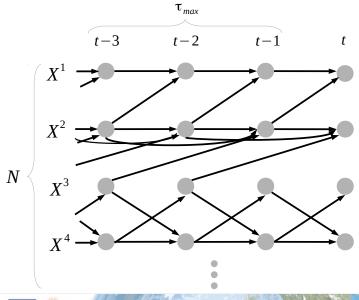






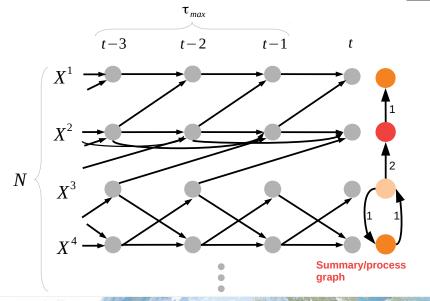






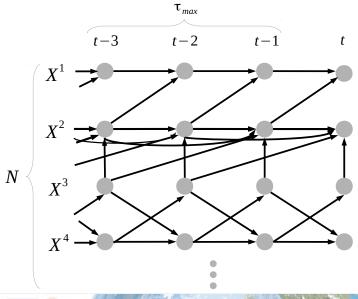






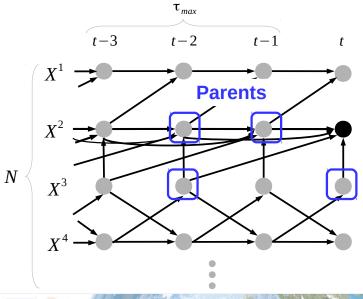














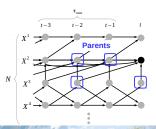


Observations in time make things easier:

• Additional constraint: Causation cannot go back in time.

Observations in time make things harder:

- High dimensionality: Resolving in time increases the number of variables
- Statistical issues: Autocorrelation makes conditional independence tests statistically harder









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Adaption of the PC-algorithm to time series in which these challenges are addressed.





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LPCMCI-Algorithm:

Generalization of PCMCI that allows for hidden confounders.





Alternative Concept: Granger Causality

Idea:

X causes Y if the past of X helps in predicting the Y from its own past.

Limitations:

- Requires causal sufficiency
- Does not allow contemporaneous interactions



State of the art: see Runge et al. NatComm 2019

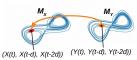
a Granger causality

$$Y_t = \sum_{\tau=1}^{p} \beta_{\tau} Y_{t-\tau} + \alpha_{\tau} X_{t-\tau} + E_t^{Y}$$
 (1)

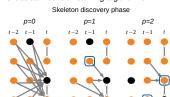
$$Y_t = \sum_{\tau=1}^{p} \beta'_{\tau} Y_{t-\tau} + E'_{t}^{Y}$$
 (2)

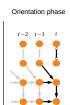
b Nonlinear state-space methods



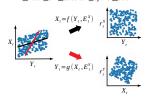


${\bf C}$ Causal network learning algorithms





d Structural causal models <u>Linear Non-Gaussian Acyclic Model</u>

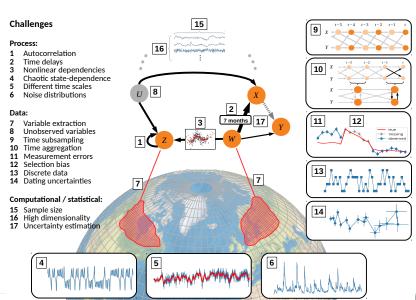


X Y

Z



Real world challenges: see Runge et al. NatComm 2019





Application examples







Application cases

Testing causal hypotheses
 [Runge et al., 2014, Runge et al., 2015b, Kretschmer et al., 2016,
 Runge et al., 2019b, Kretschmer et al., 2018, Runge et al., 2019a, Krich et al., 2020]





Application cases

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- Optimal statistical prediction schemes
 [Runge et al., 2015a, Kretschmer et al., 2017, Di Capua et al., 2019]





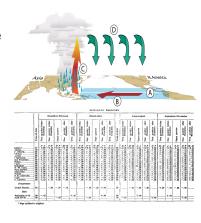
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- Evaluating climate/physical models
 [Schleussner et al., 2014, Nowack et al., 2019]



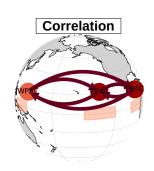


 Monthly surface pressure anomalies in the West Pacific (WPAC), surface air temperature anomalies in the Central Pacific (CPAC) and East Pacific (EPAC)





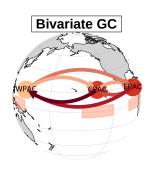
- Monthly surface pressure anomalies in the West Pacific (WPAC), surface air temperature anomalies in the Central Pacific (CPAC) and East Pacific (EPAC)
- Correlation analysis gives a completely connected graph





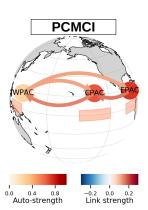


- Monthly surface pressure anomalies in the West Pacific (WPAC), surface air temperature anomalies in the Central Pacific (CPAC) and East Pacific (EPAC)
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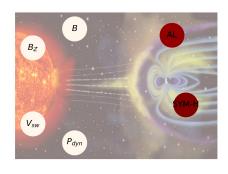


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- PCMCI [Runge et al., 2019b] better identifies the Walker circulation





 Hypothesis on interaction between magnetospheric Auroral Electrojet index (AL), magnetospheric ring current strength (SYM-H), and solar wind parameters

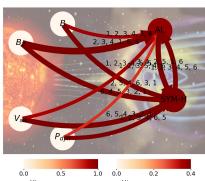


Runge et al. *Sci. Rep.* (2018), $\Delta t = 20$ min resolution



- Hypothesis on interaction between magnetospheric Auroral Electrojet index (AL), magnetospheric ring current strength (SYM-H), and solar wind parameters
- Mutual information analysis gives many dependencies

Mutual information



I^{MI}_{auto} (node color) I_{cross}^{MI} (edge color)

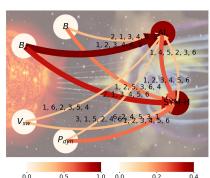
Runge et al. Sci. Rep. (2018), $\Delta t = 20 \text{min}$

resolution



- Hypothesis on interaction between magnetospheric Auroral Electrojet index (AL), magnetospheric ring current strength (SYM-H), and solar wind parameters
- Mutual information analysis gives many dependencies
- Transfer Entropy cannot remove indirect and common driver links

Transfer Entropy



I^{bivTE} (node color)

 I_{cross}^{bivTE} (edge color)

Runge et al. Sci. Rep. (2018), $\Delta t = 20$ min

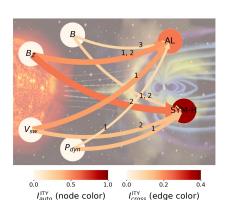
resolution



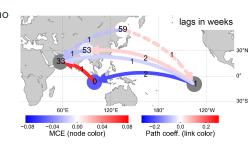
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Runge et al. *Sci. Rep.* (2018), $\Delta t = 20$ min resolution

PCMCI

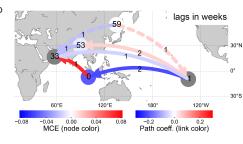


 Pathway mechanisms by which El Nino influences Indian monsoon through sea-level pressure system



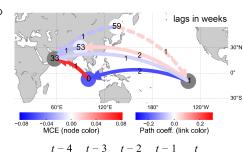


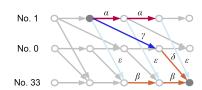
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 quantifies how much an intermediate
 variable (node) contributes to a causal
 effect





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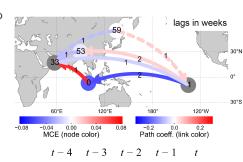


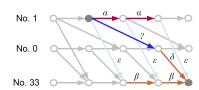






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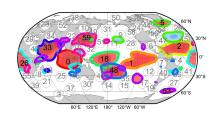






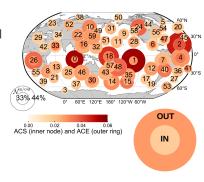


 Complex network measures based on extracted causal network from sea-level pressure system





- Complex network measures based on extracted causal network from sea-level pressure system
- Global causal gateways based on Average Causal Effect (ACE)

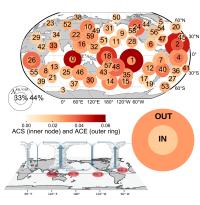








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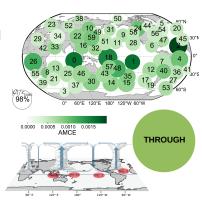








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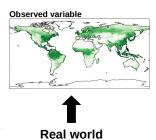




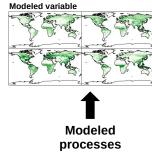




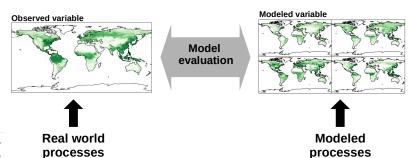
Motivation: Simple statistics (e.g. mean, variance, trends) can be right for the wrong reasons



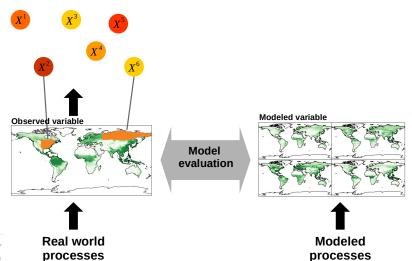
processes



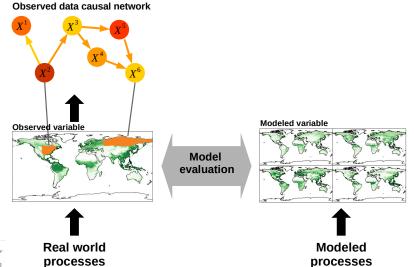
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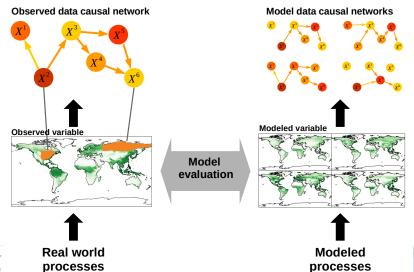
Idea: Compare climate models and observations in terms of causal relationships



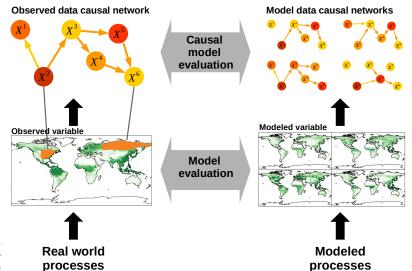
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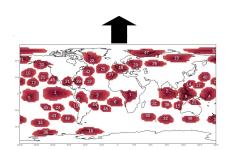


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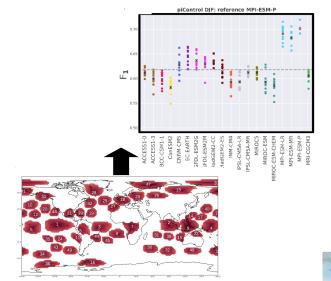


First results: CMIP5 simulations (historical and preindustrial) vs NCEP/NCAR reanalysis data of regional 3-day-mean sea level pressure



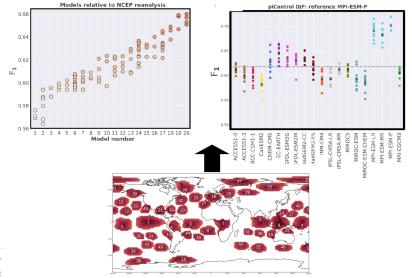


Validation: Similar climate models have similar causal networks; F-score as network comparison metric





Model evaluation: Significant differences in comparison to reanalysis





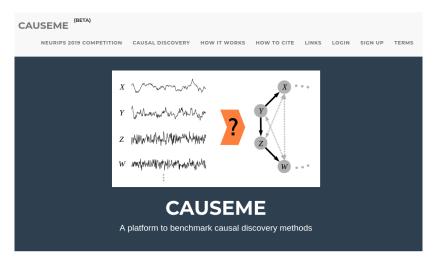
Causality benchmark platform





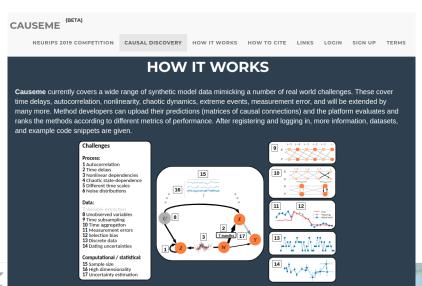


Joint work with Jordi Munoz-Mari and Gustau Camps-Valls (U Valencia)





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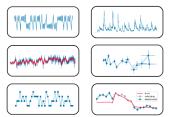
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JAKOB RUNGE

DATA AND MODELS

METHODS IFRANKING HOWTO MY RESULTS LOGOUT

DATA AND MODELS



Below you find a list of available datasets. Currently, they come from dynamical model systems featuring different challenges for causal discovery from time series as discussed in the accompanying Nature Communications Perspective paper. At the end of this page you find information on how to contribute real world datasets or model systems. Clicking on the model name will bring you to a description of the model and a list of experimental datasets. Please see the CauseMe workflow description in HowTo on how to upload your results for these experiments.

You can search through the database by name, description or tags.

DIR

Name Long name

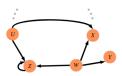
Filter models:

Joint work with Jordi Munoz-Mari and Gustau Camps-Valls (U Valencia)

JAKOB RUNGE

DATA AND MODELS METHODS IFRANKING HOWTO MY RESULTS LOGOUT

METHODS



Below you find a list of methods applied by users of this platform. Clicking on the name will bring you to a description of the method. You can search through the database by name, user, and tags. Register your own methods on My Results!



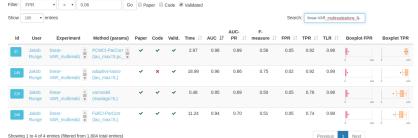


Joint work with Jordi Munoz-Mari and Gustau Camps-Valls (U Valencia)

JAKOB RUNGE DATA AND MODELS METHODS **IFRANKING** HOWTO MY RESULTS LOCOUT

RANKING

The table below presents a ranking of methods for different experiments and can be sorted according to the different metrics in columns. Optionally, the table can be filtered by metric values above or below a certain threshold. For example, one can display only methods with a FPR below 6% and sort these by TPR in decreasing order. In addition, the search field can be used on the whole table to select only particular experiments or particular methods (or both). For example, "varmodel N-10 T-150" will list all methods with 'varmodel' in the string and all experiments with N=10 variables and sample length T=150. See here for a description of metrics; AUC is based on scores, while F-measure, EPR, and TPR are based on binary link predictions by thresholding uploaded p-values at 0.05 (only available if p-values were uploaded). TLR requires lag predictions.







• Causal inference:

Answering causal questions from empirical data





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- Two settings:





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 - 1. Causal graph assumed known \rightarrow estimate causal effects





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- and to indicate how conclusions are altered for different assumptions



Thank you! Questions?

- Nature Comm. Perspective on causal discovery in time series [Runge et al., 2019a]
- Causal inference: full theory [Pearl, 2000], primer
 [Pearl et al., 2016], linear models [Pearl, 2013], popular science book
 [Pearl and Mackenzie, 2018]
- Causal discovery: general [Spirtes et al., 2000], for time series [Runge, 2018a, Runge et al., 2019a]
- Restricted SCMs [Peters et al., 2017]
- PCMCI [Runge et al., 2019b] in Science Advances
- PCMCI⁺ [Runge, 2020] in *UAI*
- LPCMCI [Gerhardus and Runge, 2020] in NeurIPS
- My software: jakobrunge.github.io/tigramite





Causal discovery in a nutshell











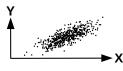










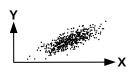










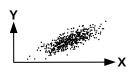










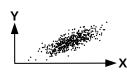














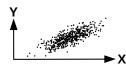




- Observed dependence: Possible causal models?
- Causal Markov Condition: statistical dependence

 ⇒ causal connectedness

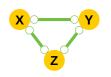


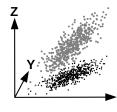






- Observed dependence: Possible causal models?
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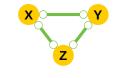


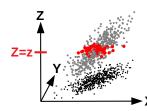






- Observed dependence: Possible causal models?
- Causal Markov Condition: statistical dependence ⇒ causal connectedness
- Observed conditional independence: $X \perp \!\!\! \perp Y \mid Z$ $\Leftrightarrow p(X, Y \mid Z) = p(X \mid Z)p(Y \mid Z)$



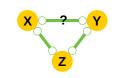


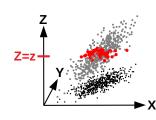






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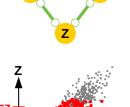


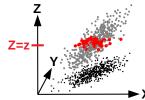






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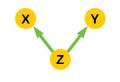


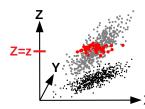






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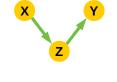


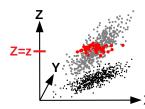






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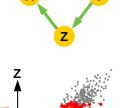


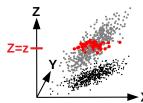






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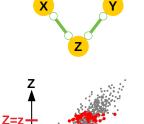








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- Markov equivalence: Cannot distinguish graphs

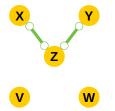








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- Suppose V⊥W, X⊥Y|Z, X⊥V|Z, X⊥W|Z, Y⊥V|Z, Y⊥W|Z, but all others are dependent (X⊥Y, X⊥W, Y⊥W, etc.)



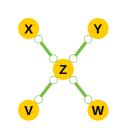








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- Suppose $V \perp \!\!\! \perp \!\!\! \perp W$, $X \perp \!\!\! \perp Y | Z$, $X \perp \!\!\! \perp V | Z$, $X \perp \!\!\! \perp W | Z$, $Y \perp \!\!\! \perp V | Z$, but all others are dependent $(X \not \!\! \perp \!\!\! \perp Y, X \not \!\! \perp \!\!\! \perp W, Y \not \!\! \perp \!\!\! \perp W$, etc.)
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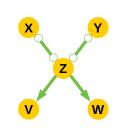








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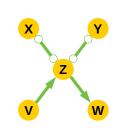








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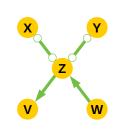








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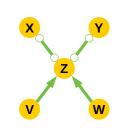








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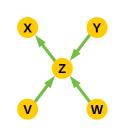








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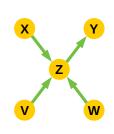








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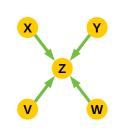








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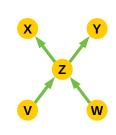








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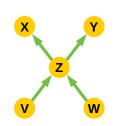








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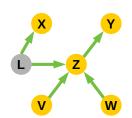








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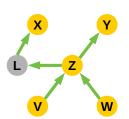








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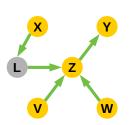








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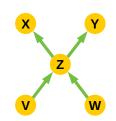








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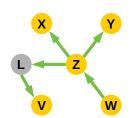








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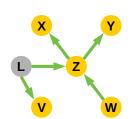








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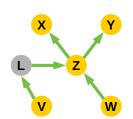








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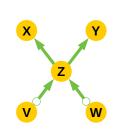








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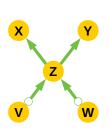


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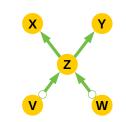








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 - Causal inference methods use (some of) these assumptions, and others (e.g., time order)















Assumptions (make causality great again)

 Bertrand Russell (1913): "The law of causality, I believe, is a relic of bygone age, surviving, like the monarchy, only because it is erroneously supposed to do no harm."









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- Assumptions are not a limitation, but a tool to arrive at causal conclusions
- Stating scientific results based on a clearly defined set of assumptions provides a basis for arguing whether a particular assumption is justified or not









Causal Sufficiency







- Causal Sufficiency
- Causal Markov Condition





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- Faithfulness





- Causal Sufficiency
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- Causal cycles





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- Assumptions on dependency types (linearity etc) and distributions
 (→ Structural Causal Models)
- Causal approaches typically utilize a subset (not all!) of these assumptions





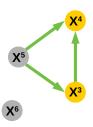








Causal Sufficiency





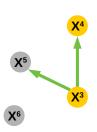








Causal Sufficiency





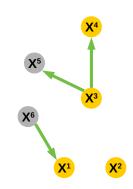








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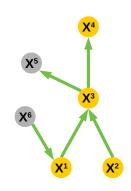








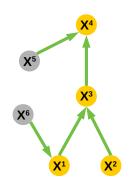
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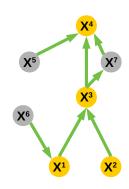
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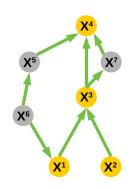
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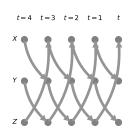
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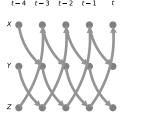
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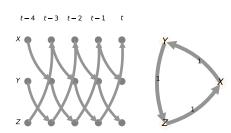


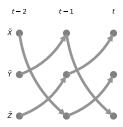






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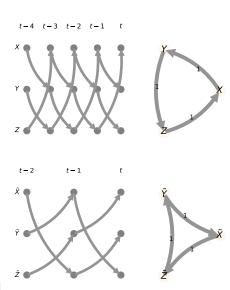








Causal Sufficiency















Pseudo-indeterministic causal graphs

A causal graph $\mathcal{G}=(\mathbf{X},\mathcal{E})$ is *pseudo-indeterministic* for a population, if and only if \mathcal{G} is <u>not</u> a deterministic causal structure for the population and there exists a causal graph \mathcal{G}' for the population over a set of variables \mathbf{X}' that properly includes \mathbf{X} such that

- 1. \mathcal{G}' is a deterministic causal structure for the population
- 2. If X and Y are in **X**, then $X \to Y$ is in $\mathcal E$ if and only if $X \to Y$ is in $\mathcal E'$
- 3. No variable in X is a cause of a variable in $X' \setminus X$
- 4. No variable in $\mathbf{X}' \setminus \mathbf{X}$ is a common cause of two variable in \mathbf{X}





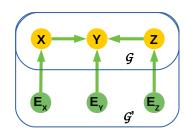




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The joint distribution $P(\mathbf{X})$ obeys Global Markov condition on \mathcal{G} iff for all disjoint subsets X^i , X^j , $S \subset \mathbf{X}$,

$$X^{i} \bowtie X^{j} \mid S \implies X^{i} \perp \!\!\!\perp X^{j} \mid S$$
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Separation implies independence and dependence implies connectedness. D-separation (for DAGs) characterizes all and only the conditional independence relations that follow from satisfying the Markov condition.

 Probability distributions for pseudo-indeterministic systems (in which the exogenous variables are independently distributed) satisfy the Markov Condition







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- Probability distributions for pseudo-indeterministic systems (in which the exogenous variables are independently distributed) satisfy the Markov Condition
- Macroscopic systems are (mostly) deterministic, but in practice we never have access to all of the causally relevant variables affecting a macroscopic system → if we include enough variables so that the excluded variables are probabilistically independent of one another, then our model will satisfy the Markov Condition



The Markov Condition holds for most systems, some arguable exeptions (plato.stanford.edu):

Artificial variables: Suppose X, Y, Z are independent and causally unrelated. Now define U = X + Y and W = Y + Z. Then U and W will be dependent, even though there is no causal relation between them.







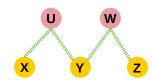






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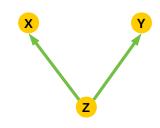






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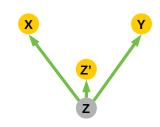






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Quantum systems: EPR

(Einstein-Podolsky-Rosen) set-up: two singlet particles are perfectly anti-correlated even if they are sufficiently far away from each other that it is impossible for one outcome to causally influence the other. Fither the Causal Markov condition or locality principle must be wrong.







N. Rosen







The Markov Condition holds for most systems, some arguable exeptions (plato.stanford.edu):

In most cases is the Markov Condition violated only for the distribution of the observed variables \rightarrow more a violation of Causal Sufficiency







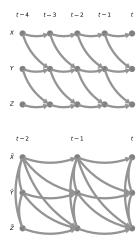
N. Rosen





The Markov Condition holds for most systems, some arguable exeptions (plato.stanford.edu):

For example, also if time series are $\ensuremath{\mathsf{aggregated}}$









Spirtes, P., C. Glymour, and R. Scheines. 2000. Causation, Prediction, and Search:

The basis for the Causal Markov Condition is, first, that it is necessarily true of populations of structurally alike pseudo-indeterministic systems whose exogenous variables are distributed independently, and second, it is supported by almost all of our experience with systems that can be put through repetitive processes and whose fundamental propensities can be tested. Any persuasive case against the Condition would have to exhibit macroscopic systems for which it fails and give some powerful reason why we should think the macroscopic natural and social systems for which we wish causal explanations also fail to satisfy the condition. It seems to us that no such case has been made.







Faithfulness







Faithfulness Condition

Definition: The joint distribution P(X) is faithful to \mathcal{G} iff every conditional independence relation true in P(X) is entailed by the Causal Markov Condition applied to \mathcal{G} .

Theorem: Faithfulness holds iff for all disjoint subsets X^i , X^j , S, $X^i \perp \!\!\! \perp X^j \mid S \implies X^i \bowtie X^j \mid S$ (2) $X^i \bowtie X^j \mid S \implies X^i \not\perp \!\!\! \perp X^j \mid S$ (contraposition)

Conditional independence implies separation and connectedness implies dependence.

• Faithfulness is violated if we observe a statistical independence relation that are not entailed by the Markov condition







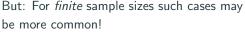
Counteracting mechanisms: For c = -ab we have $X \perp \!\!\! \perp Y$ even though they are connected in the graph:

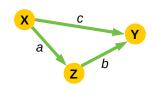
$$X = \eta^{X}$$

$$Z = aX + \eta^{Y}$$

$$Y = bZ + cX + \eta^{Z}.$$
(3)

In linear models coefficient values form real space and the set of points in this space that create vanishing partial correlations not implied by the Causal Markov Condition have Lebesgue measure zero











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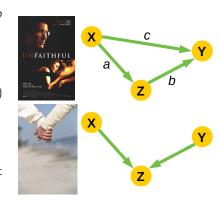
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But: For *finite* sample sizes such cases may be more common!









Determinism: Here we have

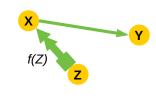
$$I(X;Y|Z) = I(f(Z);Y|Z) = 0$$
 (since $H(f(Z)|Z) = 0$) implying $X \perp \!\!\! \perp Y \mid Z$ even though Y depends on X in the model:

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$$X = f(Z)$$

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One can argue that the complexity of the underlying processes will almost always imply that variables do not deterministically depend on their parents, but some unresolved processes constitute 'intrinsic' or 'dynamical' noise.









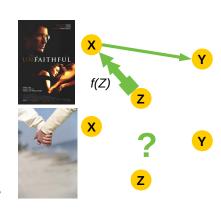
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Purely non-pairwise dependencies:

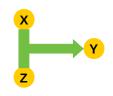
Here I(X; Y) = I(Z; Y) = 0 even though both are connected to Y:

$$Y = X \oplus Z + \eta^Y$$

X, Z binary random variables with

$$P(X = 0, 1) = P(Z = 0, 1) = 0.5$$

But if $P(X) \neq P(Z)$, Faithfulness is not violated anymore! Another way out: Hypergraphs.







Faithfulness Violations

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Faithfulness Violations

→ Non-faithful distributions arise from a pathological fine-tuning of dependence parameters

But (plato.stanford.edu):

- 'no fine-tuning' condition seems implausible as a metaphysical or conceptual constraint upon the connection between causation and probabilities (e.g., genes suppress each other)
- Faithfulness is a methodological principle rather than a metaphysical principle: it is preferable to postulate a causal structure that implies the independencies rather than one that is
 merely consistent with independence











 Here we assume directed <u>acyclic</u> graphs (DAGs) as underlying causal structure

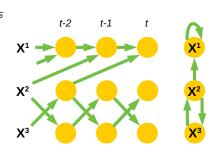








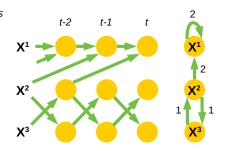
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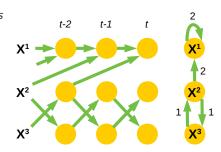
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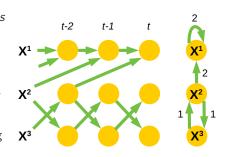








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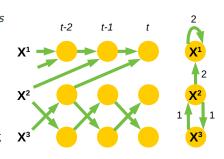








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- But what if cannot resolve time? One can think of causal graphs for systems in *equilibrium* and use the equilibrium distribution for causal discovery
- Difficult topic not further discussed



Selection bias







Selection bias

 \rightarrow covered in next lecture













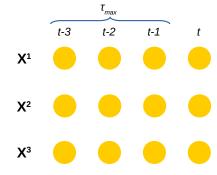




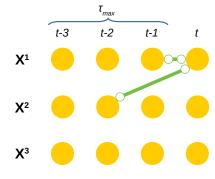




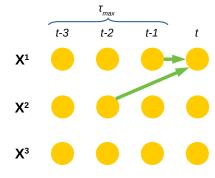






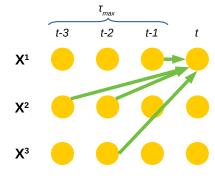






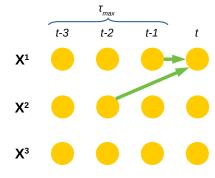






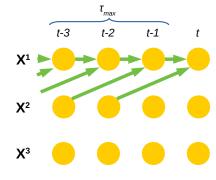






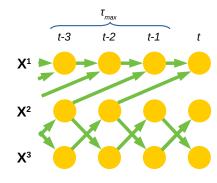








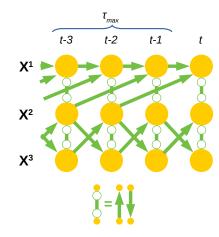








- Time order
- No contemporaneous effects: Then causal graphs can be efficiently estimated using simplifications
 → lecture on PCMCI and Granger Causality

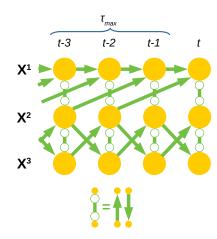








- Time order
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 → lecture on PCMCI and Granger Causality
- Stationarity

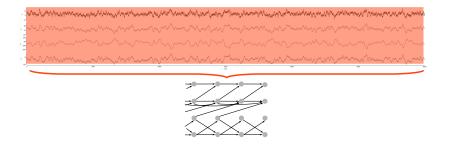








Definition: The time series process X_t with graph(s) G_t is called *causally stationary over a time index set* T iff $G_t = G_s$ for all $t, s \in T$.

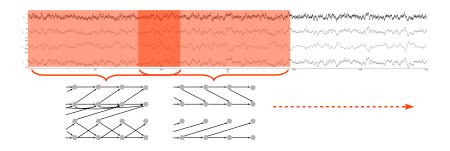








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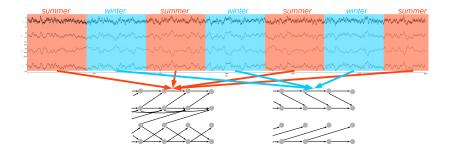








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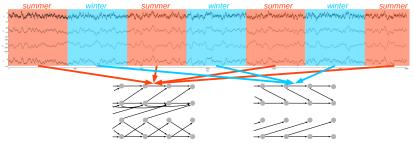






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Weaker form than the common definition of stationarity in mean, variance, spectral properties, or of the value of individual coefficients in a linear model.













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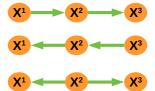






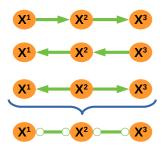


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- Conditional-independence alone can only recover causal structures up to a Markov equivalence class
- Especially useless for problems with just two variables







 Causal inference can also utilize assumptions on dependencies and distributions: E.g., linear additive models with non-gaussian noise:







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- Regression of Y on X yields residual r_Y. Suppose we measure that r_Y and X are uncorrelated, but r_X and Y are correlated



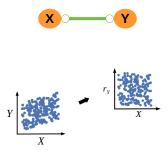








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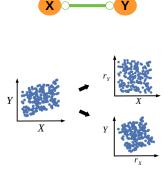








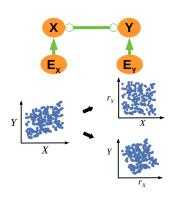
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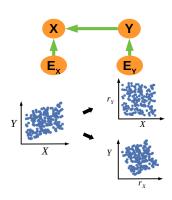








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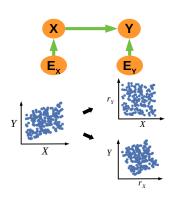








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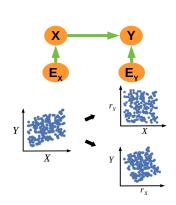






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